

# Active subspaces for Bayesian inversion

Application to a methane hydrate model

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#### Outline

1. Bayesian inversion – Basics

2. Active subspaces - General

3. Active subspaces for Bayesian inversion

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# **Bayesian** inversion

Uncertainties are modeled probabilistically.

Deterministic inverse problems are in general ill-posed.

 $\rightarrow$  Regularization by a Bayesian problem formulation.

Task: Get a probabilistic description of model parameters **x** incorporating data **d**.

$$\mathbf{d} = \mathscr{G}(\mathbf{x}^{\star}) + \eta$$
 (1)

- **d** Data
- $\mathscr{G}$  Forward/Observable map
- $\mathbf{x}^{\star}$  "True" parameter
- $\eta$  Noise



# Using Bayes' theorem

Treat parameter **x** and data **d** as random variable. Assume  $\eta \sim \mathcal{N}(0, \Gamma)$ .

Task: "Get" posterior distribution on the parameter space.

$$\rho_{\text{post}}(\mathbf{x}|\mathbf{d}) \propto \rho_{\text{like}}(\mathbf{d}|\mathbf{x})\rho_{\text{prior}}(\mathbf{x})$$
(2)

$$= \exp\left(-\frac{\|\mathbf{d} - \mathscr{G}(\mathbf{x})\|_{\Gamma}^{2}}{2}\right)\rho_{\text{prior}}(\mathbf{x})$$
(3)  
$$= \exp\left(-f_{\mathbf{d}}(\mathbf{x})\right)\rho_{\text{prior}}(\mathbf{x})$$
(4)

- $\bullet \ \|\cdot\|_{\Gamma}\coloneqq \|\Gamma^{-1/2}\cdot\|_2$
- $ho_{\text{post}}$  Posterior
- $ho_{
  m like}$  Likelihood
- $ho_{
  m prior}$  Prior
- fd Data misfit function



# Sampling from the posterior

One way: Markov chain Monte Carlo (MCMC)

Construct a Markov chain s.t. its stationary distribution is the posterior.

 $\rightarrow$  Metropolis(-Hastings) algorithm

Algorithm 1 Metropolis algorithm for the posterior.

Assume a symmetric proposal density function  $\tau$ , an initial point  $\mathbf{x}_1$ , and a prior  $\rho_{\text{prior}}$  is given. For i = 1, 2, ..., N-1

- 1. Draw a proposal  $\tilde{\mathbf{x}}$  from  $\tau$  centered at  $\mathbf{x}_i$ .
- 2. Calculate the acceptance ratio

$$\alpha(\tilde{\mathbf{x}}, \mathbf{x}_i) = \min\left(1, \frac{\exp(-f_{\mathsf{d}}(\tilde{\mathbf{x}}))\rho_{\mathsf{prior}}(\tilde{\mathbf{x}})}{\exp(-f_{\mathsf{d}}(\mathbf{x}_i))\rho_{\mathsf{prior}}(\mathbf{x}_i)}\right).$$
(5)

- 3. Draw  $u \sim \mathscr{U}([0,1])$ .
- 4. Set  $\mathbf{x}_{i+1} = \tilde{\mathbf{x}}$  if  $\alpha(\tilde{\mathbf{x}}, \mathbf{x}_i) \ge u$ , otherwise set  $\mathbf{x}_{i+1} = \mathbf{x}_i$ .



#### Mixing behavior





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# Active subspaces<sup>1</sup>

**Goal**: Approximate a function  $f : \mathbb{R}^n \to \mathbb{R}$  by a function  $g : \mathbb{R}^k \to \mathbb{R}$  (k < n), i.e. find  $\mathbf{A} \in \mathbb{R}^{n \times k}$  and g s.t.

$$f(\mathbf{x}) \approx g(\mathbf{A}^{\top}\mathbf{x}).$$
 (6)

(Also called *ridge approximation*.)

Note:  $g(\mathbf{A}^{\top}\mathbf{x})$  constant along  $\mathscr{N}(\mathbf{A}^{\top})$ .  $[\nabla(g(\mathbf{A}^{\top}\mathbf{x})) = \nabla g(\mathbf{A}^{\top}\mathbf{x})\mathbf{A}^{\top}]$ 

<sup>1</sup>[Constantine et al., 2014, Constantine, 2015] Mario Teixeira Parente (TUM) | Chair of Numerical Mathematics (M2) | Active subspaces for Bayesian inversion

# Ridge approximation

Define

$$\mathbf{C} \coloneqq \int_{\mathbb{R}^n} \nabla f \, \nabla f^\top \rho \, \, d\mathbf{x}. \tag{7}$$

C is symmetric and positive semi-definite.

 $\rightarrow \mathbf{C} = \mathbf{W} \wedge \mathbf{W}^{\top}$  for orthogonal  $\mathbf{W} \in \mathbb{R}^{n \times n}$ ,  $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_n)$ ,  $\lambda_i \ge 0$ .

$$\lambda_i = \mathbf{w}_i^{\top} \mathbf{C} \mathbf{w}_i = \int_{\mathbb{R}^n} \left( \mathbf{w}_i^{\top} \nabla f \right)^2 \rho \, d\mathbf{x}$$
(8)

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(8)



#### Eigenvalues





#### Subspaces

Decompose

$$\mathbf{W} = \begin{bmatrix} \mathbf{W}_1 & \mathbf{W}_2 \end{bmatrix} \tag{9}$$

for  $\mathbf{W}_1 \in \mathbb{R}^{n \times k}$ ,  $\mathbf{W}_2 \in \mathbb{R}^{n \times (n-k)}$  and

$$\mathbf{x} = \mathbf{W}\mathbf{W}^{\top}\mathbf{x} = \mathbf{W}_{1}\mathbf{W}_{1}^{\top}\mathbf{x} + \mathbf{W}_{2}\mathbf{W}_{2}^{\top}\mathbf{x}$$
(10)

$$= \mathbf{W}_1 \mathbf{y} + \mathbf{W}_2 \mathbf{z} \tag{11}$$

for  $y := \mathbf{W}_1^\top \mathbf{x}$ ,  $z := \mathbf{W}_2^\top \mathbf{x}$ .

**y** is called the *active*, **z** the *inactive* variable.  $\mathscr{R}(\mathbf{W}_1)$  is the *active subspace*.



# Approximation of *f*

(New) **Task**: If  $\lambda_{k+1}, \ldots, \lambda_n$  are small enough, approximate *f* by finding *g* s.t.

$$f(\mathbf{x}) \approx g(\mathbf{W}_1^{\top} \mathbf{x}). \tag{12}$$

Note:  $\mathscr{N}(\mathbf{W}_1^{\top}) = \mathscr{R}(\mathbf{W}_2)$ .

#### Algorithm 2 Computing the regression surface *g* in the active variable

Assume samples  $\mathbf{x}_i$  (i = 1, ..., N) according to  $\rho$  and corresponding function values  $f_i$  (i = 1, ..., N) are given.

1. Compute samples  $\mathbf{y}_i$  in the active subspace by

$$\mathbf{y}_i = \mathbf{W}_1^{\mathsf{T}} \mathbf{x}_i, \quad i = 1, \dots, N.$$
(13)

2. Find a regression surface g for pairs  $(\mathbf{y}_i, f_i)$  such that

$$g(\mathbf{y}_i) \approx f_i, \quad i = 1, \dots, N.$$
 (14)



# Computing $C^{[2]}$

Approximate

$$\mathbf{C} \approx \frac{1}{N} \sum_{i=1}^{N} \nabla f_i \nabla f_i^{\top} = \hat{\mathbf{W}} \hat{\Lambda} \hat{\mathbf{W}}^{\top}$$
(15)

with  $\nabla f_i \coloneqq \nabla f(\mathbf{x}_i)$  for samples  $\mathbf{x}_i \sim \rho$ .

Heuristic for *N*:

$$N = \alpha \times \ell \times \log(n)$$
 (16)

- $lpha \in [2, 10]$  oversampling factor
- $\ell$  number of eigenvalues to approximate
- *n* dimension of parameter space
- Basic tool: Eigenvalue Bernstein inequality for subexponential matrices

<sup>&</sup>lt;sup>2</sup>[Constantine and Gleich, 2014]



# Spectral gap – Theorem

#### Theorem

Let  $\varepsilon > 0$  be "small enough" and choose  $N = N(\varepsilon)$  "large enough". Then

$$dist(\mathscr{R}(\mathbf{W}_1), \mathscr{R}(\hat{\mathbf{W}}_1)) \leq \frac{4\lambda_1 \varepsilon}{\lambda_k - \lambda_{k+1}}$$
(17)

with "high probability".<sup>3</sup>

 $\left[\operatorname{dist}(\mathscr{R}(\mathbf{W}_1), \mathscr{R}(\hat{\mathbf{W}}_1)) = \|\mathbf{W}_1\mathbf{W}_1^\top - \hat{\mathbf{W}}_1\hat{\mathbf{W}}_1^\top\|_2 = \|\mathbf{W}_1^\top\hat{\mathbf{W}}_2\|_2\right]$ 

<sup>3</sup>[Constantine and Gleich, 2014]



# Bootstrap intervals and subspace distances<sup>4</sup>



(a) Eigenvalues with bootstrap intervals.

(b) Subspace distances and bootstrap intervals.

- Bootstrap: Reuse samples to get a variance-like quantity
- Compute subspace distances and also bootstrap intervals

<sup>&</sup>lt;sup>4</sup>[Constantine and Gleich, 2014]



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# Active subspaces for Bayesian inversion<sup>5</sup>

In the context of Bayesian inversion, we set

$$f(\mathbf{x}) = f_{\mathbf{d}}(\mathbf{x}) \coloneqq \frac{\|\mathbf{d} - \mathscr{G}(\mathbf{x})\|_{\Gamma}^2}{2}.$$
 (18)

Suppose we found a lower dimensional map  $g_d$  s.t.  $f_d(\mathbf{x}) \approx g_d(\mathbf{W}_1^\top \mathbf{x}) = g_d(\mathbf{y})$ . Then

$$\rho_{\text{post}}(\mathbf{x}|\mathbf{d}) \propto \rho_{\text{like}}(\mathbf{d}|\mathbf{x})\rho_{\text{prior}}(\mathbf{x})$$
(19)

$$\propto \exp(-f_{\mathsf{d}}(\mathbf{x}))\rho_{\mathsf{prior}}(\mathbf{x})$$
 (20)

$$\approx \underbrace{\exp(-g_{\mathbf{d}}(\mathbf{y}))\rho_{\text{prior}}(\mathbf{y})}_{\propto \rho_{\text{post}}(\mathbf{y}|\mathbf{d})}\rho_{\text{prior}}(\mathbf{z}|\mathbf{y})$$
(21)

 $\rightarrow$  MCMC in the active subspace.

<sup>&</sup>lt;sup>5</sup>[Constantine et al., 2016]

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# MCMC in the active subspace

Algorithm 3 Metropolis algorithm for the posterior in the active subspace

Assume a symmetric proposal density function  $\tau$ , an initial point  $\mathbf{y}_1$ , and a kernel density estimate  $\hat{\rho}$  for the prior on the active variable  $\rho_{\text{prior}}(\mathbf{y})$  is given.

For i = 1, 2, ..., N - 1

- 1. Draw a proposal  $\tilde{\mathbf{y}}$  from  $\tau$  centered at  $\mathbf{y}_i$ .
- 2. Calculate the acceptance ratio

$$\alpha(\tilde{\mathbf{y}}, \mathbf{y}_i) = \min\left(1, \frac{\exp(-g_{\mathsf{d}}(\tilde{\mathbf{y}}))\hat{\rho}(\tilde{\mathbf{y}})}{\exp(-g_{\mathsf{d}}(\mathbf{y}_i))\hat{\rho}(\mathbf{y}_i)}\right).$$
(22)

- 3. Draw  $u \sim \mathcal{U}([0,1])$ .
- 4. Set  $\mathbf{y}_{i+1} = \tilde{\mathbf{y}}$  if  $\alpha(\tilde{\mathbf{y}}, \mathbf{y}_i) \ge u$ , otherwise set  $\mathbf{y}_{i+1} = \mathbf{y}_i$ .

MCMC in the active subspace gives *active samples* **y**. Sample additionally from  $\rho_{\text{prior}}(\mathbf{z}|\mathbf{y})$  to get

$$\mathbf{x} = \mathbf{W}_1 \mathbf{y} + \mathbf{W}_2 \mathbf{z}.$$
 (23)



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### Methane hydrates

- Crystalline solids formed when water molecules enclatharate methane molecules.
- Stable at low temperatures and high pressures.
- Occur in permafrost regions and marine off-shores.
- If warmed or depressurized, they destabilize and dissociate into water and gas.
  - $\Rightarrow$  Energy content estimated to exceed the *combined* energy content of all other conventional fossil fuels<sup>6</sup>.



Figure : Burning hydrate (source: United States Geological Survey)

<sup>&</sup>lt;sup>6</sup>[Piñero et al., 2013] Mario Teixeira Parente (TUM) | Chair of Numerical Mathematics (M2) | Active subspaces for Bayesian inversion



# Methane hydrates model

- Interested in hydrates' reaction under forces/loading
  - $\rightarrow \text{Elastoplasticity model}$
- Linear relationship of stress  $\sigma$  and elastic strain  $\varepsilon^e$  through Hooke's law

$$\sigma = \mathbf{C} : \varepsilon^{e},$$

where **C** is a symmetric, positive definite, fourth order elastic stiffness tensor depending on the material (Lamé coefficients, Young's modulus, Poisson's ratio, ...).

• More details in e.g. [Gupta et al., 2015]



# Inference for model parameters<sup>8</sup>

**Goal**: Match *stress-strain curves* of the elastoplasticity model for methane hydrates and quantify uncertainties in parameters.

- Data for gas-saturated sediments from laboratory experiments<sup>7</sup>
- 8 parameters controlling evolution of curves (peak stress, hardening and softening slopes)
- Uniform prior on  $[-1,1]^8$
- 2% noise level (Γ is diagonal)



Figure :  $n_d = 46$  data points on stress and strain curves.

<sup>&</sup>lt;sup>7</sup>[Deusner et al., 2012]

<sup>&</sup>lt;sup>8</sup>[Teixeira Parente et al., 2018]



(24)

#### Computational costs

Costs lie in computing the gradient

$$abla f_{\mathsf{d}}(\mathbf{x}) = 
abla \mathscr{G}(\mathbf{x})^{ op} \Gamma^{-1}(\mathscr{G}(\mathbf{x}) - \mathbf{d})$$

with  $\nabla \mathscr{G} : \mathbb{R}^n \to \mathbb{R}^{n_{\mathbf{d}} \times n}, \mathbf{d} \in \mathbb{R}^{n_{\mathbf{d}}}.$ 

We apply (central) finite differences.  $\rightarrow$  17 forward runs

Number of gradient samples was N = 250 (trivially parallelizable).



#### Plots of eigenpairs



(a) Eigenvalues with spectral gaps after  $\lambda_2$  and  $\lambda_5$ .

#### $\rightarrow$ Comparison between 2D and 5D subspace.



(b) First 5 eigenvectors and subspaces distances.



#### 2D summary plot of fd



Figure : *r*<sup>2</sup> scores: 0.9776 (2*D*), 0.9824 (5*D*)



#### Prior/posterior on the active variable



Figure : Top row: prior on the active variable(s); middle row: posterior on the 2D active variable; bottom row: posterior on the 5D active variable



#### Posterior in 2D



Figure : Uni- and bivariate marginals of 2D posterior. Colored numbers display the absolute value of correlation coefficients.



#### Posterior in 5D



Figure : Uni- and bivariate marginals of 5D posterior. Colored numbers display the absolute value of correlation coefficients.



#### Autocorrelations + mixing





#### Posterior means



Figure : Forward evaluations with 2*D* and 5*D* posterior mean values.



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